

MATH 201 (SUMMER '23, SESS A2)

MIDTERM SOLUTIONS

2. (20 points)

Consider two urns, labelled H and T respectively. H contains two red balls. T contains five red balls, three blue balls and two green balls. Flip a coin and then uniformly draw two balls from the appropriate urn (H if heads, T if tails).

LET

R = TWO RED BALLS DRAWN

G = TWO GREEN BALLS DRAWN

B = TWO BLUE BALLS DRAWN

A = COIN-FLIP IS HEADS

T = COIN-FLIP IS TAILS

E = BOTH BALLS ARE THE
SAME COLOUR.

(a) What is the probability that you draw two green balls?

OBSERVE THAT $T = H^c$, AND SO BY THE LAW OF TOTAL PROBABILITY,

$$P(G) = P(G|H) \cdot P(H) + P(G|T) \cdot P(T)$$

RECALL $P(\cdot|H)$ & $P(\cdot|T)$ ARE UNIFORM, HENCE IS GIVEN BY

$\frac{\# \text{ FAVORABLE}}{\# \text{ TOTAL}} \Rightarrow$ ① $P(G|H) = 0$ [\because NO GREEN BALLS IN H]

② $P(G|T) = \frac{2.1}{10.9} = \frac{1}{45}$

[ASSUMING ORDER MATTERS \rightarrow CAN USE ORDER DOESN'T MATTER ALSO]

SINCE $P(H) = P(T) = \frac{1}{2}$,

$$P(G) = 0 \cdot \frac{1}{2} + \frac{1}{45} \cdot \frac{1}{2} = \frac{1}{90}$$

(b) What is the probability that you draw two red balls?

THIS IS SIMILAR, EXCEPT THAT

$$P(R | H) = 1 \quad (\because \text{BOTH BALLS IN } H \text{ ARE RED})$$

$$P(R | T) = \frac{5 \cdot 4}{10 \cdot 9} = \frac{2}{9}$$

$$\begin{aligned}\therefore P(R) &= P(R|H) P(H) + P(R|\tau) P(\tau) \\ &= 1 \cdot \frac{1}{2} + \frac{2}{9} \cdot \frac{1}{2} = \frac{11}{18}\end{aligned}$$

(c) What is the probability that you draw two blue balls?

$$\begin{aligned}\text{Ans} \quad P(B) &= P(B|H) \cdot P(H) + P(B|\tau) P(\tau) \\ &= 0 \cdot \frac{1}{2} + \frac{3 \cdot 2}{10 \cdot 9} \cdot \frac{1}{2} \\ &= \frac{1}{30}\end{aligned}$$

(d) Given that you drew two balls of the same color, what is the probability that the coin flip turned up heads?

NOTE THAT

$$E = B \cup R \cup G$$

PAIRWISE DISJOINT

$$\Rightarrow \text{BY ADDITIVITY, } P(E) = P(B) + P(R) + P(G)$$

$$= \frac{1}{30} + \frac{11}{18} + \frac{1}{90} = \frac{59}{90}$$

O.T.O.H.

$$P(E|H) = \frac{P(E \cap H)}{P(H)} = \frac{P(H)}{P(H)} = 1$$

($\because H \subseteq E \longrightarrow$ FLIPPING HEADS ALWAYS LEADS TO TWO BALLS OF THE SAME COLOUR SINCE ALL BALLS IN H ARE RED.)

$$\therefore P(H|E) = \frac{P(H \cap E)}{P(E)} = \frac{P(E|H) \cdot P(H)}{P(E)} = \frac{1 \cdot \frac{1}{2}}{\frac{59}{90}} = \frac{45}{59}$$

(BAYES' FORMULA)

3. (20 points) Hades keeps rolling a fair, standard die until he rolls a 6. Let H be the number of times he rolls the die.

(a) H is an example of a random variable you have seen before. Identify the distribution and the parameters that specify the distribution.

GEOMETRIC $\rightsquigarrow H \sim \text{Geom}(1/6)$

(b) State the general formula for the mean and variance the of random variable you identified in part (a).

IF $X \sim \text{Geom}(p)$ $(q := 1-p)$

$$E(X) = 1/p, \quad \text{Var}(X) = \frac{q}{p^2} = \frac{1-p}{q^2}$$

(c) Derive the formula for the variance you stated in part (b). (You may assume the formula for the expectation).

$$X \sim \text{Geom}(p)$$

RECALL,

$$\text{Var}(X) = E(X \cdot (X-1)) + E(X) - E(X)^2 \quad - (*)$$

$E(X) = 1/p$, so IT SUFFICES TO SHOW

$E(X \cdot (X-1)) = 2q/p^2$, AS THEIR R.H.S. OF (*) IS

$$\frac{2q}{p^2} + \frac{1}{p} - \frac{1}{p^2} = \frac{1}{p^2} [2q + p - 1] = \frac{q}{p^2} \quad (\because p+q=1)$$

NOW,

$$E(X \cdot (X-1)) = \sum_k k \cdot (k-1) P(X=k)$$
$$= \sum_{k=1}^{\infty} k \cdot (k-1) q^{k-1} p \quad (\text{p.m.f. of Geom}(p))$$

$$= \sum_{k=0}^{\infty} k \cdot (k-1) q^{k-1} p \quad (\because k=0 \text{ TERM VANISHES})$$

$$= pq \sum_{k=0}^{\infty} k \cdot (k-1) q^{k-2}$$

$$= pq \sum_{k=0}^{\infty} \frac{d^2}{dx^2} [x^k]_{x=q}$$

$$= pq \frac{d^2}{dx^2} \left[\sum_{k=0}^{\infty} x^k \right]_{x=q}$$

($\because 0 < q < 1$)

$$= pq \frac{d^2}{dx^2} \left[\frac{1}{1-x} \right]_{x=q} \quad (\text{GEOM. SERIES FORMULA})$$

$$= pq \left[\frac{2}{(1-x)^3} \right]_{x=q} = \frac{2pq}{(1-q)^3} = \frac{2pq}{p^3} = \frac{2q}{p^2} \blacksquare$$

(d) What is the mean and variance of H ?

$$p = \frac{1}{6} \Rightarrow q = 1 - p = 1 - \frac{1}{6} = \frac{5}{6}$$

$$\therefore E(H) = \frac{1}{p} = 6$$

$$\text{Var}(H) = \frac{q}{p^2} = \frac{5/6}{1/6^2} = 30$$

4. (15 points) Two events A and B are called *mutually exclusive* if $P(AB) = 0$. Suppose A , B , and C are events such that and such that both of the following are true:

- $P(A) = 0.3$, $P(B) = 0.5$, $P(C) = 0.4$.

- A and C are mutually exclusive. $\} \rightarrow P(AC) = 0$

- A and B are independent. $\} \rightarrow P(AB) = P(A)P(B) = (0.3)(0.5)$

$$= 0.15$$

Is it possible that B and C are mutually exclusive?

NO. SUPPOSE, FOR THE SAKE OF CONTRADICTION,
THAT B & C ARE MUTUALLY EXCLUSIVE.

THEN, $P(BC) = 0$.

Also

$$P(ABC) \geq 0$$

(By DEFN.)

(S.T.S.H.)

CLEARLY $ABC \subseteq AC$

$$\Rightarrow P(ABC) \leq P(AC) = 0$$

$$\Rightarrow P(ABC) = 0$$

By INCLUSION-EXCLUSION

$$P(A \cup B \cup C) = P(A) + P(B) + P(C) - P(AB) - P(BC) - P(AC) + P(ABC)$$

$$= 0.3 + 0.5 + 0.4 - 0.15 - 0 - 0 + 0 = 1.05$$

HENCE $P(A \cup B \cup C) = 1.05 > 1$

CONTRADICTION AS

PROB ≤ 1 ALWAYS

\therefore B & C CANNOT BE MUTUALLY EXCLUSIVE.

Q5 & Q6 ARE IDENTICAL TO THE
SAMPLE MIDTERM

⇒ SEE SOLNS. TO SAMPLE MIDTERM

7. (20 points) Let $Z \sim \mathcal{N}(0, 1)$ be a standard normal variable, and let $X = e^Z$. Then, X is called a log-normal random variable.

(a) Express the c.d.f. of X in terms of $\Phi(t)$. Using this, compute $P(-2 \leq X \leq 1)$.

NOTE : $X > 0$ w/ PROB. = 1 AS $X = e^Z > 0$ ALWAYS.

$$\therefore F_X(t) = 0 \quad \text{IF } t \leq 0$$

OTOH, IF $t > 0$,

$$\begin{aligned} F_X(t) &= P(X \leq t) = P(e^Z \leq t) = P(Z \leq \ln t) \\ &= \Phi(\ln t) \quad (\text{BY DEFINITION OF } \Phi) \end{aligned}$$

$$\therefore F_X(t) = \begin{cases} 0, & t \leq 0 \\ \Phi(\ln t) & t > 0 \end{cases}$$

$$\begin{aligned} \therefore P(-2 \leq X \leq 1) &= F_X(1) - F_X(-2) \\ &= \Phi(\ln 1) - 0 \\ &= \Phi(0) = 0.5 \end{aligned}$$

(b) Using your answer to part (a), explicitly compute the p.d.f. of X .

F_X IS CONTINUOUS

$\Rightarrow f_X = F_X'$ ALMOST EVERYWHERE

IF $t < 0$, $f_X'(t) = 0$

IF $t > 0$

$$f_X'(t) = \frac{d}{dt} \Phi(\ln t) \stackrel{\text{CHAIN RULE}}{=} \frac{d}{dt}(\ln t) \cdot \Phi'(\ln t)$$
$$= \frac{1}{t} \phi(\ln t) = \frac{1}{t} \left[\frac{e^{-\frac{(\ln t)^2}{2}}}{\sqrt{2\pi}} \right]$$

$$\therefore f_X(t) = \begin{cases} 0, & \text{IF } t < 0 \\ \frac{1}{t\sqrt{2\pi}} \exp\left(-\frac{(\ln t)^2}{2}\right) & \text{IF } t > 0 \end{cases}$$

(c) Compute $E(X)$. [Hint: $-\frac{1}{2}x^2 + x = -\frac{1}{2}(x - 1)^2 + \frac{1}{2}$.]

$$E(X) = \underline{F}(e^Z)$$

$$= \int_{-\infty}^{\infty} e^t \cdot \phi(t) dt$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^t \cdot e^{-t^2/2} dt = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-t^2/2 + t} dt$$

BY HINT, $-\frac{t^2}{2} + t = -\frac{1}{2}(t-1)^2 + \frac{1}{2}$

$$\begin{aligned}\therefore E(X) &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{\left(-\frac{1}{2}(t-1)^2 + \frac{1}{2}\right)} dt \\ &= e^{1/2} \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(t-1)^2} dt\end{aligned}$$

SET $u = t - 1 \Rightarrow du = dt$

$$\therefore E(X) = e^{1/2} \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-u^2/2} du$$

BUT,

$$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-u^2/2} du = \int_{-\infty}^{\infty} \phi(u) du = 1$$

($\because \phi$ is a p.d.f.)

$$\therefore E(X) = e^{1/2} \cdot 1 = e^{1/2}$$