

Sums of Poisson random variables

Proposition 17.

Let

- X_1, \dots, X_n independent random variables
- $X_i \sim \mathcal{P}(\lambda_i)$
- $Z = \sum_{i=1}^n X_i$

Then

$$Z \sim \mathcal{P}\left(\sum_{i=1}^n \lambda_i\right)$$

Consider $X_1 \sim P(\lambda_1)$, $X_2 \sim P(\lambda_2)$, $Z = X_1 + X_2$
disjoint

$$\mathbb{P}(Z=n) = \mathbb{P}(X_1 + X_2 = n) = \mathbb{P}\left(\bigcup_{k=0}^n (X_1=k, X_2=n-k)\right)$$

$$= \sum_{k=0}^n \mathbb{P}(X_1=k, X_2=n-k) \stackrel{!}{=} \sum_{k=0}^n \mathbb{P}(X_1=k) \mathbb{P}(X_2=n-k)$$

$$= \frac{1}{n!} \sum_{k=0}^n e^{-\lambda_1} \frac{\lambda_1^k}{k!} e^{-\lambda_2} \frac{\lambda_2^{n-k}}{(n-k)!} \quad n!$$

$$= \frac{e^{-(\lambda_1 + \lambda_2)}}{n!} \sum_{k=0}^n \binom{n}{k} \lambda_1^k \lambda_2^{n-k}$$

Binomial

$$\frac{e^{-(\lambda_1 + \lambda_2)}}{n!} (\lambda_1 + \lambda_2)^n$$

$$\Rightarrow Z \sim P(\lambda_1 + \lambda_2)$$

Proof for 2 random variables

Hypothesis:

$X_1 \sim \mathcal{P}(\lambda_1)$, $X_2 \sim \mathcal{P}(\lambda_2)$ and $X_1 \perp\!\!\!\perp X_2$

Computation: For $n \geq 0$,

$$\begin{aligned}\mathbf{P}(X_1 + X_2 = n) &= \sum_{k=0}^n \mathbf{P}(X_1 = k) \mathbf{P}(X_2 = n - k) \\ &= \sum_{k=0}^n e^{-\lambda_1} \frac{\lambda_1^k}{k!} e^{-\lambda_2} \frac{\lambda_2^{n-k}}{(n-k)!} \\ &= e^{-(\lambda_1 + \lambda_2)} \frac{(\lambda_1 + \lambda_2)^n}{n!}\end{aligned}$$

Sums of Binomial random variables

Proposition 18.

Let

- X_1, \dots, X_n independent random variables
- $X_i \sim \text{Bin}(n_i, p)$
- $Z = \sum_{i=1}^n X_i$

Then

$$Z \sim \text{Bin} \left(\sum_{i=1}^n n_i, p \right)$$

Outline

- 1 Joint distribution functions
- 2 Independent random variables
- 3 Sums of independent random variables
- 4 Conditional distributions: discrete case**
- 5 Conditional distributions: continuous case
- 6 Joint probability distribution of functions of random variables
- 7 Conditional expectation

General definition

Definition 19.

Let

- (X, Y) couple of discrete random variables
- Joint pmf p
- Marginal pmf's p_X, p_Y
- y such that $p_Y(y) > 0$

Then the conditional pmf of X given $Y = y$ is defined by

$$p_{X|Y}(x|y) = \mathbf{P}(\overset{\text{A}}{X = x} | \overset{\text{B}}{Y = y}) = \frac{p(x, y)}{p_Y(y)}$$

Example ctd: tossing 3 coins (1)

Experiment:

Tossing a coin 3 times

Events: We consider

$A = \text{"At most one Head"}$

$B = \text{"At least one Head and one Tail"}$

Random variables: Set

$$X_1 = \mathbf{1}_A, \quad X_2 = \mathbf{1}_B, \quad X = (X_1, X_2)$$

Example ctd: tossing 3 coins (2)

We have seen:

$X_1 \backslash X_2$	0	1	Marg. X_1
0	1/8	3/8	1/2
1	1/8	3/8	1/2
Marg. X_2	1/4	3/4	1

Conditional probabilities given $X_1 = 0$:

$$\frac{p(0,0)}{p_{X_1}(0)} = p_{X_2|X_1}(0|0) = \frac{1/8}{1/2} = \frac{1}{4}, \quad p_{X_2|X_1}(1|0) = \frac{3/8}{1/2} = \frac{3}{4}$$

$\text{Law}(X_2 | X_1=0) = \mathcal{B}(3/4) = \text{Law}(X_2) \Rightarrow$ due to $X_1 \perp\!\!\!\perp X_2$

Conditional probabilities given $X_2 = 1$:

$$\frac{p(0,1)}{p_{X_2}(1)} = p_{X_1|X_2}(0|1) = \frac{3/8}{3/4} = \frac{1}{2}, \quad p_{X_1|X_2}(1|1) = \frac{3/8}{3/4} = \frac{1}{2}$$

$\mathcal{L}(X_1 | X_2=1) = \mathcal{B}(1/2) = \mathcal{L}(X_1)$

Conditioning Poisson random variables

Proposition 20.

Let

- $X \sim \mathcal{P}(\lambda_1), Y \sim \mathcal{P}(\lambda_2)$
- $X \perp\!\!\!\perp Y$
- $p = \frac{\lambda_1}{\lambda_1 + \lambda_2}$

Then

$$\mathcal{L}(X | X + Y = n) = \text{Bin}(n, p)$$

Aim: find $\mathcal{L}(X | X+Y=n)$, that is

$$P(X=k | X+Y=n), \quad k=0, \dots, n$$

$$= \frac{P(X=k, X+Y=n)}{P(X+Y=n)}$$

$$= \frac{P(X=k, Y=n-k)}{P(X+Y=n)}$$

$$\stackrel{II}{=} \frac{P(X=k) P(Y=n-k)}{P(X+Y=n)}$$

We compute

$$X \sim P(\lambda_1) \quad Y \sim P(\lambda_2) \\ X+Y \sim P(\lambda_1 + \lambda_2)$$

$$\frac{P(X=k) P(Y=n-k)}{P(X+Y=n)}$$

$$p = \frac{\lambda_1}{\lambda_1 + \lambda_2}$$

$$= \frac{e^{-\lambda_1} \lambda_1^k e^{-\lambda_2} \lambda_2^{n-k} n!}{k! (n-k)! e^{-(\lambda_1 + \lambda_2)} (\lambda_1 + \lambda_2)^n}$$

$$= \binom{n}{k} \frac{\lambda_1^k}{(\lambda_1 + \lambda_2)^k} \frac{\lambda_2^{n-k}}{(\lambda_1 + \lambda_2)^{n-k}}$$

$$= \binom{n}{k} p^k (1-p)^{n-k}$$

$$\Rightarrow \mathcal{L}(X | X+Y=n) = \text{Bin}(n, p)$$

Proof (1)

Expression for the conditional probabilities:

Let $0 \leq k \leq n$. Then invoking $X \perp\!\!\!\perp Y$,

$$\mathbf{P}(X = k | X + Y = n) = \frac{\mathbf{P}(X = k) \mathbf{P}(Y = n - k)}{\mathbf{P}(X + Y = n)}$$

Law of $X + Y$: We have seen

$$X + Y \sim \mathcal{P}(\lambda_1 + \lambda_2)$$

Proof (2)

Computation of the conditional probabilities:

$$\begin{aligned} \mathbf{P}(X = k | X + Y = n) &= e^{-\lambda_1} \frac{\lambda_1^k}{k!} e^{-\lambda_2} \frac{\lambda_2^{n-k}}{(n-k)!} \left[e^{-(\lambda_1 + \lambda_2)} \frac{(\lambda_1 + \lambda_2)^n}{n!} \right]^{-1} \\ &= \binom{n}{k} p^k (1-p)^{n-k} \end{aligned}$$

Conclusion:

$$\mathcal{L}(X | X + Y = n) = \text{Bin}(n, p)$$

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General definition

Definition 21.

Let

- (X, Y) couple of continuous random variables
- Joint density f
- Marginal densities f_X, f_Y
- y such that $f_Y(y) > 0$

Then the conditional density of X given $Y = y$ is defined by

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)}$$

Justification of the definition

Heuristics: $f_{X|Y}(x|y)$ can be interpreted as

$$\begin{aligned} f_{X|Y}(x|y) dx &= \frac{f(x, y) dx dy}{f_Y(y) dy} \\ &\approx \frac{\mathbf{P}(x \leq X \leq x + dx, y \leq Y \leq y + dy)}{\mathbf{P}(y \leq Y \leq y + dy)} \\ &= \mathbf{P}(x \leq X \leq x + dx | y \leq Y \leq y + dy) \end{aligned}$$

Use of the conditional probability: compute probabilities like

$$\mathbf{P}(X \in A | Y = y) = \int_A f_{X|Y}(x|y) dx$$

Rigorous definition: see MA 539