

Outline

- 1 Joint distribution functions
- 2 Independent random variables
- 3 Sums of independent random variables
- 4 Conditional distributions: discrete case
- 5 Conditional distributions: continuous case**
- 6 Joint probability distribution of functions of random variables
- 7 Conditional expectation

General definition

Definition 21.

Let

- (X, Y) couple of continuous random variables
- Joint density f
- Marginal densities f_X, f_Y
- y such that $f_Y(y) > 0$

Then the conditional density of X given $Y = y$ is defined by

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)}$$

Justification of the definition

Heuristics: $f_{X|Y}(x|y)$ can be interpreted as

$$\begin{aligned} f_{X|Y}(x|y) dx &= \frac{f(x, y) dx dy}{f_Y(y) dy} \\ &\approx \frac{\mathbf{P}(x \leq X \leq x + dx, y \leq Y \leq y + dy)}{\mathbf{P}(y \leq Y \leq y + dy)} \\ &= \mathbf{P}(x \leq X \leq x + dx | y \leq Y \leq y + dy) \end{aligned}$$

Use of the conditional probability: compute probabilities like

$$\mathbf{P}(X \in A | Y = y) = \int_A f_{X|Y}(x|y) dx$$

Rigorous definition: see MA 539

Simple example of continuous conditioning (1)

Density: Let (X, Y) be a random vector with density

$$f(x, y) = \frac{e^{-\frac{x}{y}} e^{-y}}{y} \mathbf{1}_{(0, \infty)}(x) \mathbf{1}_{(0, \infty)}(y)$$

Question: Compute

$$\mathbf{P}(X > 1 | Y = y)$$

Strategy: (i) Compute $f_Y(y)$

(ii) Compute $f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)}$

(iii) Compute $\mathbf{P}(X > 1 | Y = y) = \int_1^{\infty} f_{X|Y}(x|y) dx$

$$f(x,y) = \frac{e^{-x/y} e^{-y}}{y} \mathbb{1}_{(0,\infty)}(x) \mathbb{1}_{(0,\infty)}(y)$$

Step (i) Compute

$\int_{\mathbb{R}} f(x) dx = 1$
for f density

$$f_Y(y) = \int_{\mathbb{R}} f(x,y) dx$$

$$= \int_{\mathbb{R}} \frac{e^{-x/y} e^{-y}}{y} \mathbb{1}_{(0,\infty)}(x) \mathbb{1}_{(0,\infty)}(y) dx$$

$$= e^{-y} \mathbb{1}_{(0,\infty)}(y) \int_0^{\infty} \frac{1}{y} e^{-x/y} dx$$

Density of r.v $E(\frac{1}{y})$

$$\Rightarrow f_Y(y) = e^{-y} \mathbb{1}_{(0,\infty)}(y)$$

Remark $Y \sim E(1)$

Recalling the main densities

<u>Law</u>	<u>Density</u>
$U(a, b)$	$\frac{1}{b-a} \mathbb{1}_{(a, b)}(x)$
$E(\lambda)$	$\lambda e^{-\lambda x} \mathbb{1}_{(0, \infty)}(x)$
$N(0, 1)$	$\frac{1}{\sqrt{2\pi}} e^{-x^2/2}$
$N(\mu, \sigma^2)$	$\frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$

$$f(x,y) = \frac{e^{-xy} e^{-y}}{y} \mathbb{1}_{(0,\infty)}(x) \mathbb{1}_{(0,\infty)}(y)$$

Step (ii) Compute $f_{X|Y}(x|y)$, if $x, y > 0$

$$f_{X|Y}(x|y) = \frac{f(x,y)}{f_Y(y)} = \frac{e^{-xy} e^{-y}}{y e^{-y}}$$

$$= \begin{cases} \frac{e^{-xy}}{y} & \text{if } x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

Rmk $\mathcal{L}(X|Y=y) = \mathcal{E}\left(\frac{1}{y}\right)$

Step (iii) We wish to compute

$$\mathbb{P}(X > 1 \mid Y = y)$$

$$= \int_1^{\infty} f_{X|Y}(x|y) dx$$

$$= \int_1^{\infty} \frac{1}{y} e^{-x/y} dx$$

$$= \mathbb{P}(Z > 1), \text{ with } Z \sim \mathcal{E}\left(\frac{1}{y}\right)$$

$$= e^{-1/y} \quad (\text{Formula for tails of } \mathcal{E}\left(\frac{1}{y}\right))$$

Otherwise: $\mathbb{P}(X > 1 \mid Y = y)$

$$= -e^{-x/y} \Big|_1^{\infty} = e^{-1/y}$$

Simple example of continuous conditioning (2)

Marginal distribution of Y : We have

$$\begin{aligned}f_Y(y) &= \int_0^{\infty} f(x, y) dx \\&= \frac{e^{-y}}{y} \left(\int_0^{\infty} e^{-\frac{x}{y}} dx \right) \mathbf{1}_{(0, \infty)}(y) \\&= e^{-y} \mathbf{1}_{(0, \infty)}(y)\end{aligned}$$

Conditional density: For $y > 0$ we have

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \frac{e^{-\frac{x}{y}}}{y} \mathbf{1}_{(0, \infty)}(x)$$

Namely $\mathcal{L}(X|Y = y) = \mathcal{E}\left(\frac{1}{y}\right)$

Simple example of continuous conditioning (3)

Conditional probability:

$$\begin{aligned}\mathbf{P}(X > 1 | Y = y) &= \int_1^{\infty} f_{X|Y}(x|y) dx \\ &= \int_1^{\infty} \frac{e^{-x/y}}{y} dx \\ &= e^{-\frac{1}{y}}\end{aligned}$$

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Characterizing r.v by expected values

Notation:

$C_b(\mathbb{R}^2) \equiv$ set of continuous and bounded functions on \mathbb{R}^2 .

Theorem 22.

Let $X = (X_1, X_2)$ be a r.v in \mathbb{R}^2 . We assume that

$$\mathbf{E}[\varphi(X_1, X_2)] = \int_{\mathbb{R}^2} \varphi(x_1, x_2) f(x_1, x_2) dx_1 dx_2,$$

for all functions $\varphi \in C_b(\mathbb{R}^2)$.

Then (X_1, X_2) is continuous, with density f .

Application: change of variable

Problem: Let

- $X = (X_1, X_2)$ random variable with density f .
- Set $Y = h(X)$ with $h : \mathbb{R}^2 \rightarrow \mathbb{R}^2$.

We wish to find the density of Y .

Application: change of variable (2)

Recipe: One proceeds as follows

- 1 For $\varphi \in C_b(\mathbb{R}^2)$, write

$$\mathbf{E}[\varphi(Y)] = \mathbf{E}[\varphi(h(X))] = \int_{\mathbb{R}^2} \varphi(h(x_1, x_2)) f(x_1, x_2) dx_1 dx_2.$$

- 2 Change variables $y = h(x)$ in the integral.
After some elementary computations we get

$$\mathbf{E}[\varphi(Y)] = \int_{\mathbb{R}^2} \varphi(y_1, y_2) g(y_1, y_2) dy_1 dy_2.$$

- 3 This characterizes Y , which admits a density g

Polar coordinates of Gaussian vectors (1)

Standard Gaussian vector in \mathbb{R}^2 : Consider

- $X, Y \sim \mathcal{N}(0, 1)$, with $X \perp\!\!\!\perp Y$
- $Z = (X, Y)$

Polar coordinates: Set

$$(X, Y) = (R \cos(\Theta), R \sin(\Theta))$$

Question:

Find the joint density of (R, Θ)

Model

$$X \sim W(0,1) \Rightarrow f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

$$Y \sim W(0,1) \Rightarrow f_Y(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}$$

$$X \perp\!\!\!\perp Y$$

$\Rightarrow Z = (X, Y)$ has density

$$f(x, y) = f_X(x) f_Y(y)$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \times \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}$$

$$= \frac{1}{2\pi} \exp\left(-\frac{1}{2}(x^2 + y^2)\right)$$

$$f(x,y) = \frac{1}{2\pi} \exp\left(-\frac{1}{2}(x^2+y^2)\right)$$

General formula

We let $(x = r \cos(\theta), y = r \sin(\theta)) = g^{-1}(r, \theta)$
Then the density in (r, θ) is

$$h(r, \theta) = f(g^{-1}(r, \theta)) J(r, \theta) \mathbb{1}_{(r, \theta) = g(x, y)}$$

Here

$$\cos^2(\theta) + \sin^2(\theta) = 1$$

$$\begin{aligned} \bullet f(g^{-1}(r, \theta)) &= \frac{1}{2\pi} \exp\left(-\frac{1}{2}(r^2 \cos^2 \theta + r^2 \sin^2 \theta)\right) \\ &= \frac{1}{2\pi} e^{-\frac{1}{2}r^2} \end{aligned}$$

$$\bullet \quad x = (r \cos(\theta), r \sin(\theta))$$

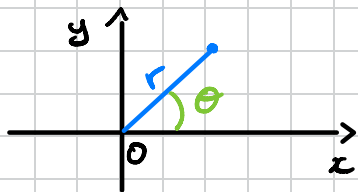
$$\Rightarrow J(r, \theta) = \left| \det \begin{pmatrix} \frac{\partial x}{\partial r} & \frac{\partial x}{\partial \theta} \\ \frac{\partial y}{\partial r} & \frac{\partial y}{\partial \theta} \end{pmatrix} \right|$$

$$= \begin{vmatrix} \cos(\theta) & -r \sin(\theta) \\ \sin(\theta) & r \cos(\theta) \end{vmatrix}$$

$$= r (\cos^2(\theta) + \sin^2(\theta))$$

$$= \boxed{r = J(r, \theta)}$$

• Recall $x = r \cos(\theta)$ $y = r \sin(\theta)$



We get $r > 0$

$$\theta \in (0, 2\pi)$$

Conclusion . The density for (R, Θ) is

$$h(r, \theta) = \frac{1}{2\pi} e^{-\frac{r^2}{2}} \times r \times \mathbb{1}_{(0, 2\pi)}(\theta) \mathbb{1}_{(0, \infty)}(r)$$

Polar coordinates of Gaussian vectors (2)

Decomposition of the expected value: For $\varphi \in \mathcal{C}_b(\mathbb{R}^2)$,

$$\begin{aligned}\mathbf{E}[\varphi(R, \Theta)] &= \mathbf{E}[\varphi(R, \Theta) \mathbf{1}_{(Y>0)}] + \mathbf{E}[\varphi(R, \Theta) \mathbf{1}_{(Y<0)}] \\ &\equiv A_+ + A_-\end{aligned}$$

Expression for A_+ :

$$\begin{aligned}A_+ &= \mathbf{E}\left[\varphi\left((X^2 + Y^2)^{1/2}, \tan^{-1}\left(\frac{Y}{X}\right)\right) \mathbf{1}_{(Y>0)}\right] \\ &= \int_{\mathbb{R} \times \mathbb{R}_+} \varphi\left((x^2 + y^2)^{1/2}, \tan^{-1}\left(\frac{y}{x}\right)\right) \frac{e^{-\frac{x^2+y^2}{2}}}{2\pi} dx dy\end{aligned}$$

Polar coordinates of Gaussian vectors (3)

Change of variable for A_+ : Set

$$x = r \cos(\theta), \quad y = r \sin(\theta) \quad \implies \quad J(r, \theta) = r$$

Then

$$A_+ = \int_{\mathbb{R}_+ \times (0, \pi)} \varphi(r, \theta) \frac{r e^{-\frac{r^2}{2}}}{2\pi} dr d\theta$$

Change of variable for A_- : We find

$$A_- = \int_{\mathbb{R}_+ \times (\pi, 2\pi)} \varphi(r, \theta) \frac{r e^{-\frac{r^2}{2}}}{2\pi} dr d\theta$$

Polar coordinates of Gaussian vectors (4)

Expression for the expected value:

$$\mathbf{E}[\varphi(R, \Theta)] = \int_{\mathbb{R}_+ \times (0, 2\pi)} \varphi(r, \theta) \frac{r e^{-\frac{r^2}{2}}}{2\pi} dr d\theta$$

Joint density for (R, Θ) :

$$f(r, \theta) = \frac{1}{2\pi} \mathbf{1}_{(0, 2\pi)}(\theta) \times r e^{-\frac{r^2}{2}} \mathbf{1}_{\mathbb{R}_+}(r)$$

Otherwise stated:

- $R \sim \text{Rayleigh}$, $\Theta \sim \mathcal{U}([0, 2\pi])$
- $R \perp\!\!\!\perp \Theta$

Change of variable: general result (repeated)

Theorem 23.

Let

- $X = (X_1, X_2)$ continuous random variable
- Density: $f_X \rightarrow$ joint density
- g diffeomorphism of \mathbb{R}^2
- $Y = g(X)$

Then Y has a density f_Y given by

$$f_Y(y) = f_X(g^{-1}(y)) J(y) \mathbf{1}_{\{y=g(x) \text{ for some } x\}}$$