

Rmk In general

X, Y decorrelated $\not\Rightarrow X \perp Y$

i.e. if X, Y are centered

$$E[XY] = 0 \not\Rightarrow X \perp Y$$

However if (X, Y) is a Gauss. vector

X, Y decorrelated $\Rightarrow X \perp Y$

Another way to state this property:

(X, Y) has a covariance matrix
which is diagonal

If (X, Y) Gauss
vector

$$\Rightarrow X \perp Y$$

$$\text{cov}(X, Y) = \begin{pmatrix} V(X) & E[XY] \\ E[XY] & V(Y) \end{pmatrix}$$

Example of (X, Y) with

$$E[XY] = 0, \text{ but } X \not\perp Y$$

Take

$$X \sim N(0, 1) \quad \varepsilon \perp\!\!\!\perp X \quad P(\varepsilon = \pm 1) = \frac{1}{2}$$

$$Y = \varepsilon X$$

$$\begin{aligned} \text{Then } E[XY] &= E[\varepsilon X^2] \\ &= E[\varepsilon] E[X^2] = 0 \end{aligned}$$

But $X \not\perp Y \rightarrow$ check $P(X \in A, Y \in B) \neq P(X \in A) P(Y \in B)$

Back to our example: we have seen that

$$\mathbb{E}[(X - \alpha Y) Y] = 0 \quad (\text{decorrelation})$$

In addition $(X - \alpha Y, Y)$ is a Gauss. vector

$$\text{Thus } (X - \alpha Y) \perp Y$$

$$\Rightarrow (X - \alpha Y) \perp \psi(Y)$$

$$\Rightarrow \mathbb{E}[(X - \alpha Y) \psi(Y)] = 0 \quad \leftarrow \text{Hyp: } \psi(Y) \in \mathcal{L}^2$$

$$\Rightarrow \alpha Y = \mathbb{E}[X | Y]$$

Rmk

If X, Y are centered, $X \perp Y$, then

$$E[XY] \stackrel{!}{=} E[X] E[Y] = 0$$

Proof

Step 1: We look for α such that

$$Z = X - \alpha Y \implies Z \perp\!\!\!\perp Y.$$

Recall: If (Z, Y) is a Gaussian vector
 $\iff Z \perp\!\!\!\perp Y$ iff $\text{cov}(Z, Y) = 0$

Application: $\text{cov}(Z, Y) = \mathbf{E}[Z Y]$. Thus

$$\text{cov}(Z, Y) = \mathbf{E}[(X - \alpha Y) Y] = \mathbf{E}[X Y] - \alpha V(Y),$$

et

$$\text{cov}(Z, Y) = 0 \quad \text{iff} \quad \alpha = \frac{\mathbf{E}[XY]}{V(Y)}.$$

Proof (2)

Step 2: We invoke (i) in the definition of π .

\hookrightarrow Let $V \in L^2(\sigma(Y))$. Then

$$Y \perp\!\!\!\perp (X - \alpha Y) \implies V \perp\!\!\!\perp (X - \alpha Y)$$

and

$$\mathbf{E}[(X - \alpha Y) V] = \mathbf{E}[X - \alpha Y] \mathbf{E}[V] = 0.$$

Thus

$$\alpha Y = \pi_{\sigma(Y)}(X) = \mathbf{E}[X | Y].$$

Outline

- 1 Definition
 - Baby conditional distributions: discrete case
 - Baby conditional distributions: continuous case
 - Definition with measure theory
- 2 Examples
- 3 Existence and uniqueness
- 4 Conditional expectation: properties
- 5 Conditional expectation as a projection
- 6 **Conditional regular laws**
 - Probability laws and expectations
 - Definition of the CRL

Aim of this section

Recall: We have seen that if

- $X \sim \mathcal{P}(\lambda_1), Y \sim \mathcal{P}(\lambda_2)$
- $X \perp\!\!\!\perp Y$
- $p = \frac{\lambda_1}{\lambda_1 + \lambda_2}$,

then

$$\mathcal{L}(X | X + Y = n) = \text{Bin}(n, p)$$

Question:

How to translate this

↔ to the non-baby conditional expectation language?

Rmk In the Poisson case

$$\mathcal{L}(X | \overbrace{X+Y}^S = n) = \text{Bin}(n, p)$$

translation $\rightarrow \mathcal{L}(X | S) = \text{Bin}(S(w), p)$

2 Problems

(i) We have computed $\mathbb{E}[\varphi(X) | S]$

\rightarrow Find a way to go from \mathbb{E} to \mathcal{L}

(ii) What is a random probability measure?

Outline

- 1 Definition
 - Baby conditional distributions: discrete case
 - Baby conditional distributions: continuous case
 - Definition with measure theory
- 2 Examples
- 3 Existence and uniqueness
- 4 Conditional expectation: properties
- 5 Conditional expectation as a projection
- 6 **Conditional regular laws**
 - **Probability laws and expectations**
 - Definition of the CRL

Characterizing r.v by expected values

Notation:

$C_b(\mathbb{R}) \equiv$ set of continuous and bounded functions on \mathbb{R} .

Theorem 25.

Let X be a r.v. We assume that

$$\mathbf{E}[\varphi(X)] = \int_{\mathbb{R}} \varphi(x) f(x) dx, \quad \text{for all functions } \varphi \in C_b(\mathbb{R}).$$

Then X is continuous, with density f .

Here, we characterize $\mathcal{L}(X)$ through
 $\{ \mathbf{E}[\varphi(X)]; \varphi \in C_b(\mathbb{R}) \}$

Application: change of variable

Standard method: $\mathbb{P}(Y \leq y) = \mathbb{P}(h(X) \leq y)$

\rightarrow we get the cdf of X (F_X)
then $f_Y = F'_Y$

Problem: Let

- X random variable with density f .
- Set $Y = h(X)$ with $h : \mathbb{R} \rightarrow \mathbb{R}$.

We wish to find the density of Y .

Application: change of variable (2)

Recipe: One proceeds as follows

$Y = h(X)$
 X has density f

- 1 For $\varphi \in C_b(\mathbb{R})$, write

$$\mathbf{E}[\varphi(Y)] = \mathbf{E}[\varphi(h(X))] = \int_{\mathbb{R}} \varphi(\underbrace{h(x)}_{=y}) f(x) dx.$$

- 2 Change variables $y = h(x)$ in the integral.

After some elementary computations we get

$$\mathbf{E}[\varphi(Y)] = \int_{\mathbb{R}} \varphi(y) g(y) dy.$$

- 3 This characterizes Y , which admits a density g

Example: normal r.v and linear transformations

Proposition 26.

Let

- $X \sim \mathcal{N}(0, 1)$
- $\mu \in \mathbb{R}$ and $\sigma > 0$
- Set $Y = \sigma X + \mu$

Then

$$Y \sim \mathcal{N}(\mu, \sigma^2)$$

Proof $X \sim W(0,1)$, $Y = \sigma X + \mu$.

Take $\varphi \in C_b(\mathbb{R})$. Then

$$\begin{aligned} \mathbb{E}[\varphi(Y)] &= \mathbb{E}[\varphi(\sigma X + \mu)] \\ &= \int_{\mathbb{R}} \varphi(\sigma x + \mu) \frac{e^{-\frac{x^2}{2}}}{\sqrt{2\pi}} dx \end{aligned}$$

cv: $y = \sigma x + \mu \Rightarrow x = \frac{y - \mu}{\sigma} \quad dx = \frac{dy}{\sigma}$

$$= \int_{\mathbb{R}} \varphi(y) \frac{e^{-\frac{(y-\mu)^2}{2\sigma^2}}}{\sqrt{2\pi\sigma^2}} dy \quad \forall \varphi \in C_b(\mathbb{R})$$

\Rightarrow density of Y : $\frac{e^{-\frac{y-\mu}{2\sigma^2}}}{(2\pi\sigma^2)^{\frac{1}{2}}} \Rightarrow \boxed{Y \sim W(\mu, \sigma^2)}$

Proof

Recipe, item 1: for $\varphi \in C_b(\mathbb{R})$, write

$$\mathbf{E}[\varphi(Y)] = \mathbf{E}[\varphi(\sigma X + \mu)] = \int_{\mathbb{R}} \varphi(\sigma x + \mu) \frac{e^{-x^2/2}}{\sqrt{2\pi}} dx.$$

Recipe, item 2: Change of variable: $y = \sigma x + \mu$:

$$\mathbf{E}[\varphi(Y)] = \int_{\mathbb{R}} \varphi(y) g(y) dx, \quad \text{with} \quad g(y) = \frac{e^{-(y-\mu)^2/(2\sigma^2)}}{\sqrt{2\pi\sigma^2}}.$$

Recipe, item 3:

Y is continuous with density g , therefore $Y \sim \mathcal{N}(\mu, \sigma^2)$.

Characterizing r.v by expected values (ctd)

Theorem 27.

Let $X : \Omega \rightarrow \mathbb{R}$ be a r.v. Then

$\{\mathbf{E}[\varphi(X)]; \varphi \in C_b(\mathbb{R})\}$ characterizes the law of X

$$\{\mathbf{E}[e^{iux}]; u \in \mathbb{R}\}$$

Outline

- 1 Definition
 - Baby conditional distributions: discrete case
 - Baby conditional distributions: continuous case
 - Definition with measure theory
- 2 Examples
- 3 Existence and uniqueness
- 4 Conditional expectation: properties
- 5 Conditional expectation as a projection
- 6 Conditional regular laws**
 - Probability laws and expectations
 - Definition of the CRL

Definition 28.

Let

- (Ω, \mathcal{F}, P) a probability space
- (S, \mathcal{S}) a measurable space of the form $\mathbb{R}^d, \mathbb{Z}^d$
- $X : (\Omega, \mathcal{F}) \rightarrow (S, \mathcal{S})$ a random variable in $L^1(\Omega)$
- \mathcal{G} a σ -algebra such that $\mathcal{G} \subset \mathcal{F}$.

We say that $\mu : \Omega \times \mathcal{S} \rightarrow [0, 1]$ is a **Conditional regular law** of X given \mathcal{G} if

- For all $f \in C_b(S)$, the map $\omega \mapsto \mu(\omega, f)$ is a random variable, equal to $\mathbf{E}[f(X) | \mathcal{G}]$ a.s.
- ω -a.s. $f \mapsto \mu(\omega, f)$ is a probability measure on (S, \mathcal{S}) .

Discrete example

Poisson law case: Let

- $X \sim \mathcal{P}(\lambda)$ and $Y \sim \mathcal{P}(\mu)$
- $X \perp\!\!\!\perp Y$

We set $S = X + Y$.

Then

CRL of X given S is $\text{Bin}(S, p)$, with $p = \frac{\lambda}{\lambda + \mu}$

Proof: we know that

$$\mathcal{L}(X | S=n) = \text{Bin}(n, p)$$

We have also seen (for $\varphi \in \mathcal{G}_b(\mathbb{R})$)

$$\begin{aligned} & \mathbb{E}[\varphi(X) | S] \quad g = \sigma(S) \\ &= \sum_{k=0}^S \binom{S}{k} p^k (1-p)^{S-k} \end{aligned}$$

Claim 1: (i) in the definition \tilde{w} satisfied, i.e. for a fixed φ ,

$\mathbb{E}[\varphi(X) | S]$ is a r.v.

Justification

$$S: \Omega \rightarrow \mathbb{R} \text{ measurable} \\ \omega \mapsto S(\omega)$$

$$\mathbb{E}[\varphi(x) | S]$$

$$= \sum_{k=0}^S \binom{S}{k} p^k (1-p)^{S-k}$$

$$= \sum_{k=0}^{\infty} \binom{S}{k} p^k \underbrace{(1-p)^{S-k}}_{\text{exp function}} \underbrace{\mathbb{1}(S \geq k)}_{\text{indicator function}}$$

polynomial
in S

exp
function

indicator
function

This is a r.v (measurable)

Justification of (ii). If we fix ω ,
we have that $S = \sum_{i=1}^n \mathbb{1}_{\{A_i\}}(\omega)$

$$S(\omega) \in \mathbb{N}$$

and $\mathcal{L}(X | S)(\omega)$

is $\text{Bin}(S(\omega), p)$

This is a probability distribution

Conclusion: we can write safely

$$\mathcal{L}(X | S) = \text{Bin}(S, p)$$

Proof for the discrete example

Proof: we have seen that for $n \leq m$

$$\mathbf{P}(X = n | S = m) = \binom{m}{n} p^n (1 - p)^{m-n} \quad \text{with } p = \frac{\lambda}{\lambda + \mu}.$$

Then we consider

- State space = \mathbb{N} , $\mathcal{G} = \sigma(S)$

and we verify that **these conditional probabilities define a CRL.**