

Randomized Coordinate Descent

$$\min_{x \in \mathbb{R}^n} f(x)$$

$$\nabla f(x)^{(i)} = \begin{bmatrix} 0 \\ \vdots \\ 0 \\ \nabla f(x)_i \\ 0 \\ \vdots \end{bmatrix}$$

① Gradient Descent is $x_{k+1} = x_k - \eta \nabla f(x_k)$

② Coordinate Descent $x_{k+1} = x_k - \eta \nabla f(x_k^{i(k)})$

$$i(1) = 1$$

$$i(2) = 2$$

only $i(k)$ -th entry of x_k is updated

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③ Randomized Coordinate Descent

$$x_{k+1} = x_k - \eta \nabla f(x_k^{i(k)})$$

$i(k) \sim \text{i.i.d. uniform distribution in } \{1, 2, \dots, n\}$

identical
independent
distributed

$$\text{Prob}(i(k)=1) = \frac{1}{n}$$

Consider an operator $T: \mathbb{R}^n \rightarrow \mathbb{R}^n$
 $x \mapsto T(x)$

and a fixed point iteration

$$x_{k+1} = T(x_k)$$

① An operator S is nonexpansive if

$$\|S(x) - S(y)\| \leq \|x - y\|$$

Example : If $\nabla f(x)$ is L-cont. and $f(x)$ is convex

then $S(x) = [I - \frac{2}{L} \nabla f](x)$ is nonexpansive

$$\begin{aligned}\|S(x) - S(y)\|^2 &= \|-\frac{2}{L}[\nabla f(x) - \nabla f(y)] + (x - y)\|^2 \\ &= \|x - y\|^2 + \frac{4}{L^2} \|\nabla f(x) - \nabla f(y)\|^2 - \frac{4}{L} \langle x - y, \nabla f(x) - \nabla f(y) \rangle \\ &\leq \|x - y\|^2\end{aligned}$$

Proximal Point Method
 $x_{k+1} = T(x_k)$

2.2.3 Convergence for convex functions

$$T = \text{Prox}_f^\eta = (I + \eta \nabla f)^{-1}$$

Theorem 2.8. Assume $\nabla f(x)$ is Lipschitz-continuous with Lipschitz constant L and $f(x) : \mathbb{R}^n \rightarrow \mathbb{R}$ is convex. Then for any x, y :

1. $f(y) \geq f(x) + \langle \nabla f(x), y - x \rangle + \frac{1}{2L} \|\nabla f(x) - \nabla f(y)\|^2$
2. $\|\nabla f(x) - \nabla f(y)\|^2 \leq L \langle \nabla f(x) - \nabla f(y), x - y \rangle.$

② $T = (1-\theta)I + \theta S$ with $\theta \in (0, 1)$

is called θ -averaged if S is nonexpansive

Example : $S = [I - \frac{2}{L} \nabla f]$

$$T = I - \eta \nabla f = (1-\theta)I + \theta(I - \frac{2}{L} \nabla f)$$

$$\theta = \frac{\eta L}{2} \in (0, 1) \iff 0 < \eta < \frac{2}{L}$$

③ If S is nonexpansive with at least one fixed point

$x_{k+1} = S(x_k)$ may not converge to x_*

Example: $S(x) = -x$ $x_* = 0$

Theorem If $S: \mathbb{R}^n \rightarrow \mathbb{R}^n$ is nonexpansive, then

$$x_{k+1} = (1-\theta)x_k + \theta S(x_k), \quad 0 < \theta < 1$$

converges to one fixed point of $S(x)$.

Example: This implies GD converges if $\eta < \frac{2}{L}$.

$$(4) \quad T(x) = \begin{bmatrix} [T(x)]_1 \\ [T(x)]_2 \\ \vdots \\ [T(x)]_n \end{bmatrix} \quad T_i(x) = \begin{bmatrix} x_1 \\ \vdots \\ x_{i-1} \\ [T(x)]_i \\ x_{i+1} \\ \vdots \\ x_n \end{bmatrix}$$

If $x_{k+1} = x_k - \eta \nabla f(x_k) = T(x_k)$ is GD

$$x_{k+1} = x_k - \eta \nabla f(x)^{i(k)} \Leftrightarrow x_{k+1} = T_{i(k)}(x_k)$$

Theorem Assume

- ① T is θ -averaged with at least one fixed point
- ② $i(k) \in \{1, \dots, n\}$ is i.i.d. with uniform probability.

$$\text{Then } x_{k+1} = T_{i(k)}(x_k)$$

converges to one fixed point of $T(x)$ with probability 1.

Example: $S = I - \frac{2}{L} \nabla f$ is nonexpansive if $f(x)$ is convex
 $\Rightarrow T = I - \eta \nabla f$
 $= (1-\theta)I + \theta S$ is θ -averaged
if $\theta = \frac{\eta L}{2} < 1 \Leftrightarrow \eta < \frac{2}{L}$

Proof: Define R, R_i by $\begin{cases} T = I - \theta R \\ T_i = I - \theta R_i \end{cases}$

$$\Rightarrow R_i(x) = \begin{bmatrix} 0 \\ \vdots \\ 0 \\ [R(x)]_i \\ \vdots \\ 0 \end{bmatrix} \quad R = \frac{1}{\theta} \nabla f = \frac{2}{L} \nabla f$$

$$T = I - \theta \cdot \frac{2}{L} \nabla f$$

$$x_{k+1} = T_{i(k)}(x_k) \Leftrightarrow x_{k+1} = x_k - \theta R_{i(k)}[x_k]$$

T is θ -averaged $\Leftrightarrow T = (1-\theta)I + \theta S$
with S nonexpansive

$$I - R = I - \frac{I - T}{\theta} \Leftrightarrow \frac{1}{\theta} T - (\frac{1}{\theta} - 1)I \text{ is nonexpansive}$$

$$= (1 - \frac{1}{\theta})I + \frac{1}{\theta} T \Leftrightarrow I - R \text{ is nonexpansive}$$

$$= S \Leftrightarrow \|x - Rx - y + Ry\|^2 \leq \|x - y\|^2$$

$$\Leftrightarrow \frac{1}{\epsilon} \|Rx - Ry\|^2 \leq \langle x - y, Rx - Ry \rangle$$

$$T(x_*) = x_* \Leftrightarrow R(x_*) = 0$$

$$y = x_* \Rightarrow \frac{1}{\epsilon} \|Rx\|^2 \leq \langle Rx, x - x_* \rangle$$

$$\frac{1}{\epsilon} \|\frac{2}{L} \nabla f(x)\|^2 \leq \langle \frac{2}{L} \nabla f(x), x - x_* \rangle$$

$$\Leftrightarrow \|\nabla f(x)\|^2 \leq L \langle \nabla f(x), x - x_* \rangle$$

Definition of Expectation & more :

Example: ① X is a random variable taking values in $\{0, 1\}$ with equal probability

definition

$$\left\{ \begin{array}{l} E(X) = 0 \cdot \frac{1}{2} + 1 \cdot \frac{1}{2} = \frac{1}{2} \\ E(X^2) = 0^2 \cdot \frac{1}{2} + 1^2 \cdot \frac{1}{2} = \frac{1}{2} \\ E(f(X)) = f(0) \cdot \frac{1}{2} + f(1) \cdot \frac{1}{2} = \frac{1}{2}f(1) \end{array} \right.$$
$$P(X=0) = \frac{1}{2}$$
$$P(X=1) = \frac{1}{2}$$

E denotes expectation w.r.t. random variable X

Expectation for discrete random variables are defined as above

② X is a random variable taking values

$$\text{in } \{x_1, x_2, \dots, x_k\}$$

with probability P_1, P_2, \dots, P_k $\sum_{i=1}^k P_i = 1, P_i \geq 0$

definition

$$\left\{ \begin{array}{l} E(X) = x_1 \cdot P_1 + x_2 \cdot P_2 + \dots + x_k \cdot P_k \rightarrow \text{convex combination} \\ E(f(X)) = \sum_{i=1}^k f(x_i) P_i \\ \text{Prob}(X=x_i) = P_i \end{array} \right.$$

Jensen's: $f(E(X)) \leq E(f(X))$ if $f(x)$ is convex

Example: $f(x) = x^2 \Rightarrow [E(X)]^2 \leq E(X^2)$

③ X, Y are i.i.d. random variable taking values

$$\{a_1, a_2, \dots, a_k\}$$

with probability P_1, P_2, \dots, P_k

$$\sum_{i=1}^k P_i = 1, P_i \geq 0$$

Joint probability $P(X=a_i, Y=a_j)$

X & Y are independent \Rightarrow

$$P(X=a_i, Y=a_j) = P(X=a_i) P(Y=a_j)$$

$$\begin{aligned} E[f(X, Y)] &= \sum_{i=1}^k \sum_{j=1}^k f(a_i, a_j) P(X=a_i, Y=a_j) \\ &= \sum_{i=1}^k \sum_{j=1}^k f(a_i, a_j) P_i P_j \end{aligned}$$

In $X_{k+1} = T_{i(k)}(X_k)$

$$X_{k+1} = X_k - \eta \nabla f(X_k)^{i(k)}$$

X_0 is deterministic and $i(0), i(1), \dots, i(N)$ are random

X_N is a function of these $N+1$ random variables

$E(X_N)$ denotes expectation w.r.t. $N+1$ i.i.d. random variables

④ Conditional Probability & Expectation

X is a random variable taking values

In $\{x_1, x_2, \dots, x_k\}$

with probability P_1, P_2, \dots, P_k

Y is a random variable taking values

In $\{y_1, y_2, \dots, y_\ell\}$

with probability q_1, q_2, \dots, q_ℓ

Probability of event $X=x_i$ given the knowledge $Y=y_j$

$$P(X=x_i | Y=y_j) = \frac{P(X=x_i, Y=y_j)}{P(Y=y_j)} \leftarrow \text{joint prob.}$$

The conditional expectation w.r.t. X given $Y=y_j$

$$\begin{aligned} E(X | Y=y_j) &= \sum_{i=1}^k x_i P(X=x_i | Y=y_j) \\ &= \sum_{i=1}^k x_i \frac{P(X=x_i, Y=y_j)}{P(Y=y_j)} \end{aligned}$$

$$g(y) = E(X | Y=y) = \sum_{i=1}^k x_i P(X=x_i | Y=y)$$

is $\begin{cases} \text{a function of } y \\ \text{a random variable} \end{cases}$

We can also write it as $g(Y) = E(X|Y)$

So $E(X|Y)$ is a random variable.

$$E_k = E[i(k) | i(k-1), i(k-2), \dots, i(0)]$$

E_k denotes the conditional expectation w.r.t. $i(k)$

Conditioned on the past random variables

$$i(k-1), i(k-2), \dots, i(0)$$

$$X_{k+1} = T_{i(k)}(X_k)$$

$$X_{k+1} = X_k - \eta \nabla f(X_k) i(k)$$

Then ① $E_k(X_k) = X_k$ because X_k does
↓
random variable NOT depend on $i(k)$

② Let X be something depending on all $i(k)$
 $i(k-1), i(k-2), \dots, i(0)$

$$E[E_k(X)] = E[X]$$

Law of Total Expectation : $E[E(X|Y)] = E(X)$

$E(X|Y)$ is a function of Y (random variable)

$$E(X|Y) = \sum_{i=1}^k x_i P(X=x_i | Y)$$

$$\begin{aligned} E[E(X|Y)] &= E\left[\sum_{i=1}^k x_i P(X=x_i | Y)\right] \\ &= \sum_{j=1}^l \left[\sum_{i=1}^k x_i P(X=x_i | Y=y_j) \right] \cdot P(Y=y_j) \\ &= \sum_{j=1}^l \sum_{i=1}^k x_i P(X=x_i | Y=y_j) \cdot P(Y=y_j) \\ &= \sum_{j=1}^l \sum_{i=1}^k x_i P(X=x_i, Y=y_j) \\ &= \sum_{i=1}^k x_i \left[\sum_{j=1}^l P(X=x_i, Y=y_j) \right] \end{aligned}$$

$$= \sum_{i=1}^k x_i P(X=x_i) \\ = E(X)$$

③ $E_k [R_{i(k)}(X_k)] = \frac{1}{n} R(X_k)$

$$R = \begin{matrix} \text{def} \\ \sum \end{matrix} \quad R_i(x) = \sum \begin{pmatrix} 0 \\ \vdots \\ 0 \\ \nabla f(x)_i \\ \vdots \\ 0 \end{pmatrix}$$

E_k is expectation w.r.t. $i(k)$

$$\left\{ \begin{array}{l} P(i(k)=1) = \frac{1}{n} \\ P(i(k)=2) = \frac{1}{n} \\ \vdots \\ P(i(k)=n) = \frac{1}{n} \end{array} \right.$$

$$\Rightarrow E_k [R_{i(k)}(X_k)] = \sum_{j=1}^n P[i(k)=j] \cdot R_j(X_k) \\ = \frac{1}{n} \sum_{j=1}^n R_j(X_k) = \frac{1}{n} R(X_k)$$

$$E_k \|R_{i(k)}(X_k)\|^2 = \frac{1}{n} \|R(X_k)\|^2$$

$$\|R_{i(k)}(X_k)\|^2 = \left([R(X_k)]_{i(k)} \right)^2$$

$$E_k \left([R(X_k)]_{i(k)} \right)^2 = \sum_{j=1}^n P[i(k)=j] \cdot (R_j(X_k))^2$$

$$= \frac{1}{n} \sum_{j=1}^n (R_j(x_k))^2 = \frac{1}{n} \|R(x_k)\|^2$$

④ Let x_* be a fixed point to $x_* = T(x_*)$

$$\begin{aligned}\|x_{k+1} - x_*\|^2 &= \|x_k - \theta R_{i(k)}(x_k) - x_*\|^2 \\ &= \|x_k - x_*\|^2 - 2\theta \langle R_{i(k)}(x_k), x_k - x_* \rangle \\ &\quad + \theta^2 \|R_{i(k)}(x_k)\|^2\end{aligned}$$

Take E_k (conditional expectation w.r.t. $i(k)$)

Expectation is linear

$$\begin{aligned}E_k \|x_{k+1} - x_*\|^2 &= E_k \|x_k - x_*\|^2 - 2\theta E_k \langle R_{i(k)}(x_k), x_k - x_* \rangle \\ &\quad + \theta^2 E_k \|R_{i(k)}(x_k)\|^2 \\ &= \|x_k - x_*\|^2 - 2\theta \langle E_k [R_{i(k)}(x_k)], x_k - x_* \rangle \\ &\quad + \theta^2 \cdot \frac{1}{n} \|R(x_k)\|^2 \\ &= \|x_k - x_*\|^2 - \frac{2\theta}{n} \langle R(x_k), x_k - x_* \rangle \\ &\quad + \theta^2 \cdot \frac{1}{n} \|R(x_k)\|^2\end{aligned}$$

$$\rightarrow \leq \|x_k - x_*\|^2 - \frac{\theta}{n} \|R(x_k)\|^2 + \frac{\theta^2}{n} \|R(x_k)\|^2$$

$$\boxed{\frac{1}{2} \|R(x)\|^2 \leq \langle R(x), x - x_* \rangle}$$

Take E for both sides

$$\Rightarrow E \|x_{k+1} - x_*\|^2 \leq E \|x_k - x_*\|^2 - (1-\theta) \frac{\theta}{n} E \|R(x_k)\|^2$$

$$\Rightarrow \{E\|X_k - X_*\|\} \downarrow$$

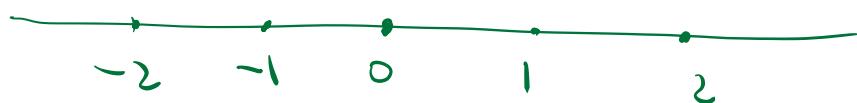
⑤ Discrete Martingale

is a sequence of random variable X_1, X_2, X_3, \dots satisfying

$$① E(|X_n|) < +\infty$$

$$② E(X_{n+1} | X_1, \dots, X_n) = X_n$$

Example : positions of 1D random walk



Integer grid ; $X_0 = 0$; $X_{k+1} = \begin{cases} X_k + 1 & \text{with prob. } \frac{1}{2} \\ X_k - 1 & \text{with prob. } \frac{1}{2} \end{cases}$

$$E(|X_n|) \leq n$$

Given X_1, \dots, X_n , expected value of X_{n+1} is X_n

$$E(X_{n+1} | X_1, \dots, X_n) = X_n$$

Supermartingale if $E(X_{n+1} | X_1, \dots, X_n) \leq X_n$

Submartingale if $E(X_{n+1} | X_1, \dots, X_n) \geq X_n$

Supermartingale Convergence Theorem

Let X_k, Y_k be random variables satisfying

① $X_k \geq 0, Y_k \geq 0$ almost surely

② $E[X_{k+1} | x_1, \dots, x_k] \leq X_k - Y_k$

Then almost surely (with probability 1)

1) X_k converges to X_∞ , $k \rightarrow \infty$

2) $\sum_{k=0}^{\infty} Y_k < +\infty$

Remark: X_∞ is a random variable

$$E_k \|X_{k+1} - X_*\|^2 \leq \|X_k - X_*\|^2 - (1-\theta) \frac{\theta}{n} \|R(X_k)\|^2$$

$$\Rightarrow \left\{ \sum_{k=0}^{\infty} \|R(X_k)\|^2 < \infty \Rightarrow \|R(X_k)\| \rightarrow 0 \right.$$

$\lim_{k \rightarrow \infty} \|X_k - X_*\|^2$ exists with probability 1

Let $\text{Fix}(T)$ be the set of all fixed points of T

$\forall x_* \in \text{Fix}(T) \Rightarrow \left[\lim_{k \rightarrow \infty} \|X_k - X_*\| \text{ exists with prob. 1} \right]$

 With prob. 1 $\left[\forall x_* \in \text{Fix}(T), \lim_{k \rightarrow \infty} \|X_k - X_*\| \text{ exists} \right]$

Nontrivial, skipped, see last reference book.

⑥ With prob. 1, we have

1) $\{X_k\}$ is a bounded sequence

thus a convergent subsequence $\{X_{k_j}\} \rightarrow z$

$$\begin{aligned} \Rightarrow \|R(x_k)\| \rightarrow 0 &\Rightarrow \|\frac{1}{\theta}(I-T)(x_k)\| \rightarrow 0 \\ &\Rightarrow \|(I-T)x_k\| \rightarrow 0 \end{aligned}$$

T is θ -averaged

$$\Rightarrow \|(I-T)x_{k_j}\| \rightarrow 0$$

$$\Rightarrow \|T(x) - T(y)\| \leq \|x - y\|$$

$$\Rightarrow \|z - T(z)\| = 0$$

T is continuous

$$\Rightarrow z - T(z) = 0$$

$I - T$ is continuous

$$\Rightarrow z \in \text{Fix}(T)$$

$$\begin{aligned} \|x_{k+1} - x_*\|^2 &= \|x_k - x_*\|^2 - 2\theta \langle R(x_k) x_k, x_k - x_* \rangle \\ &\quad + \theta^2 \|R(x_k)\|^2 \end{aligned}$$

$$\boxed{\|R(x_k)\| \rightarrow 0}$$

set $x_* = z \rightarrow \{x_k\}$ has the same limit as x_{k_j}